

# Downtime-Frequency Curves for Availability Characterization

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## Abstract

*Availability data typically consist of start and end times of a system's down periods. We propose a natural way to plot their statistics so that Service Level Agreements concerning availability can be formulated as the condition that an empirical curve lies below a given curve.*

## 1. Introduction

In this paper we have in mind data communications systems, whose dependability was recently considered in VTT's IPLU project [2], but most of the discussion applies more generally. The traditional way to set a quantitative requirement for the availability of a system is to give a single number like 0.99999. Such numbers are typically used in Service Level Agreements (SLA) concerning, for example, transmission links. A single-number characteristic is, however, quite uninformative: it tells nothing about the lengths of the individual downtimes, which may have great significance.

## 2. Definition and properties

We propose instead the use of *downtime-frequency curves* that characterize the frequency of each down-period length separately in an appropriate form. They are defined as follows.

Consider first the characterization of the reliability of a system or, similarly, availability of a resource, with binary nature: at each timepoint  $t$ , it can be unequivocally stated whether the system is up or down. Thus, its performance is described by a  $\{0, 1\}$ -valued stochastic process:

$$I_t = 1_{\{\text{system down at time } t\}}.$$

The probability of failure,  $\mathbb{P}(\text{system down at time } t) = \mathbb{E}I_t$  is already a characteristic of the reliability of the system.

Assuming stationarity and ergodicity, this number is independent of  $t$  and obtained almost surely as the limit of the observed relative frequency:

$$\lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T I_t dt = \mathbb{E}I_0 \quad \text{a.s.}$$

Let us define the ongoing down-period length at time  $t$  as

$$W_t = \inf\{s \geq t : I_s = 0\} - \sup\{s \leq t : I_s = 0\}.$$

When the system is up, we have  $W_t = 0$ . The relative share of time spent in down-periods lasting longer than  $\tau$  during an observation period of length  $T$  is then given by the random variable

$$\varphi_T(\tau) = \frac{1}{T} \int_0^T 1_{\{W_t > \tau\}} dt.$$

Considered as a random function of  $\tau$ ,  $\varphi_T(\tau)$  is non-increasing. Its initial value  $\varphi_T(0)$  equals the relative overall downtime of the system in the observation period (for example one year).

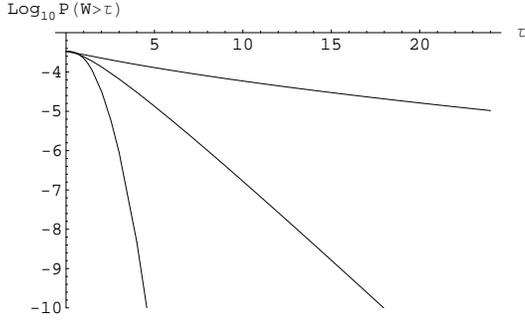
If the system is stationary,  $W_t$  is a stationary stochastic process, and we find that the expectation of the random function  $\varphi_T(\tau)$  equals the tail distribution function of the random variable  $W = W_0$ :

$$F_T(\tau) = \mathbb{E} \varphi_T(\tau) = \frac{1}{T} \int_0^T \mathbb{P}(W_t > \tau) dt = P(W > \tau).$$

Using this framework, we can now formulate reliability criteria that take into account the down-period lengths also: let us consider the performance of the system acceptable if

$$\varphi_T(\tau) \leq \psi(\tau)$$

for some selected function  $\psi$ . The function  $\psi$  can be specified in a SLA. The network operator has to build the system in such a way that the expected curve  $F_T(\tau)$  lies sufficiently much below  $\psi(\tau)$ . Since the relevant values of both the down-periods and the probabilities extend over many orders of magnitude, the curves should be drawn in a log-plot or even log-log plot. When the axes are selected appropriately, the curve  $\varphi_T(\tau)$  can often be given as a straight line.



**Figure 1. The expected downtime-frequency curves when the downtimes are Weibull distributed with exponents  $\alpha = 0.5$  (highest curve), 1 (middle curve), 2 (lowest curve).**

### 3 On/off-availability and quality-availability

The binary notion of availability is often insufficient for communication systems. The packet transfer may work ‘reliably’ in both directions but proceed with much lower rate and/or higher delays than in normal conditions. From a mathematical point of view, however, this problem can be reduced to the binary case simply by considering the *set* of binary processes

$$1_{\{q(S_t) \leq r\}}, \quad r \in R,$$

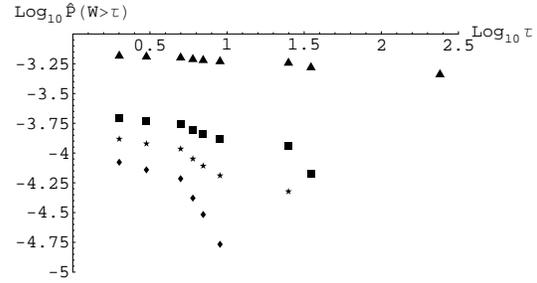
where  $S_t$  is the system state at time  $t$ ,  $q$  is some characteristic of it (rate, delay, . . .), and  $R$  is the set of possible or relevant values of that characteristic.

For example, an SLA may require that the bandwidth of an MPLS path be higher than 50 Mbit/s with an availability of at least 0.999. Then, the set  $R$  may contain the value 50 Mbit/s alone. However, since IP-based services are usually quite flexible with respect to bandwidth requirements, it would make sense to require additionally that the availability of 5 Mbit/s be at least 0.99999.

One can also, at least in principle, let  $R$  be a whole interval and replace the binary-case criterion that the empirical values should lie below a curve to the two-dimensional criterion that they should lie below a surface. If higher  $q(\cdot)$  means better quality, the monotonicities behave similarly in both dimensions if  $r$  is replaced by some inverse parameter  $\beta$  by writing, for example,  $r = 1/\beta$ .

### 4. Examples

A standard mathematical model of this kind of process is the alternating renewal process, where the up- and down-periods are independent random variables with distributions



**Figure 2. Data plot example (see text).**

$G_{\text{up}}$  and  $G_{\text{down}}$  and means  $\mu_{\text{up}}$  and  $\mu_{\text{down}}$ , respectively. (The resulting formulae can in fact be generalized using the Palm theory of stationary processes, see [1].) When  $I_t$  is a stationary version on an alternating renewal process, the distribution of  $W$  is

$$\mathbb{P}(W > \tau) = \frac{1}{\mu_{\text{up}} + \mu_{\text{down}}} \int_{\tau}^{\infty} y G_{\text{down}}(dy).$$

Note that the distribution  $G_{\text{up}}$  has an effect only through the expectation  $\mu_{\text{up}}$ .

Here is a formal example of such plots. Assume that the down-periods and up-periods are independent, time unit is one hour, the up- and down-periods have means 3000 and 1, respectively, and the down-period length has a Weibull distribution

$$1 - G_{\text{down}}(y) = \exp(-\beta y^{\alpha}),$$

where  $\alpha$  and  $\beta$  are parameters. The choice  $\beta = \Gamma(1 + 1/\alpha)^{\alpha}$  yields the desired mean 1. We can now compute and plot the functions  $F_T(\tau)$  for three qualitatively different parameter values  $\alpha = 0.5, 1$  and  $2$ . This example also illustrates the usability of linear, log-linear and log-log plots for various purposes.

As an example how empirical data might look like in this framework, assume that the downtimes of a system within a year consist of intervals with lengths 2, 2, 2, 3, 3, 5, 5, 6, 7, 9, 25, 35, 240 minutes (in ascending order). The empirical tail distribution function of  $W$  is then determined by the points marked as triangles in Figure 2. Note that the few long down-periods have the effect that the whole point set looks almost horizontal. The other three point sets show the corresponding plot when 1, 2 and 3 largest values are removed from the data set, respectively.

### References

- [1] F. Baccelli and P. Bremaud. *Elements of Queueing Theory*. Springer Verlag, Berlin, 2003.
- [2] IPLU project homepage. <http://iplu.vtt.fi>.